

# statistics cheat sheet

Nasser M. Abbasi

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**problem:** phone calls received at rate  $\lambda = 2$  per hr. If person wants to take 10 min shower, what is probability a phone will ring during that time?

answer: first change to  $\omega = \lambda \frac{10}{60} = 2 \frac{10}{60} = .3333$ , now we want  $P(X \geq 1) = 1 - P(X \leq 1) = 1 - P(0)$

but  $P(k) = \frac{\lambda^k}{k!} e^{-\lambda}$ , but remember, we are using  $\omega$ , so  $P(k) = \frac{\omega^k}{k!} e^{-\omega}$  so  $P(0) = \frac{.3333^0}{0!} e^{-.3333} = 0.777$   
so  $P(X \geq 1) = 1 - .777 = 0.283$ , so 28% change the phone will ring.

How long can shower be if they wish probability of receiving no phone calls to be at most 0.5?

$P(0) = 0.5 = \frac{\omega^0}{0!} e^{-\omega} \rightarrow 0.5 = e^{-\omega}$  hence  $\ln 0.5 = -\omega \rightarrow \omega = 0.693$ , so  $\lambda \frac{x}{60} = 0.693 \rightarrow x = 20.7$  minutes

**To find quantile**, say  $\frac{1}{4}$ , first find an expression for  $F(x)$  as function of  $x$ , then solve for  $x$  in  $F(x) = .25$

For median, solve for  $x$  in  $F(x) = .5$

**properties of CDF:** 1. Show  $F(x) \geq 0$  for all  $x$ . Do this by showing  $F'(x) \geq 0$ , and show limit  $F(x) \rightarrow 1$  as  $x \rightarrow \infty$  and limit  $F(x) \rightarrow 0$  as  $x \rightarrow -\infty$ . And  $P(k_1 \leq T < k_2) = F(k_2) - F(k_1)$

**properties of pdf:**

1. piecewise continuous
2.  $\text{pdf}(x) \geq 0$
3.  $\int_{-\infty}^{\infty} \text{pdf}(x) = 1$

**remember**  $\frac{d}{dx} \tan^{-1} x = \frac{1}{1+x^2}$

The geometric distribution is the only discrete memoryless random distribution. It is a discrete analog of the exponential distribution. continuous

Some relations

$$\sum_{k=1}^n k = \frac{1}{2} n(n+1)$$

Geometric sum

$$\sum_{k=0}^n r^k = \frac{1 - r^{n+1}}{1 - r}$$

if  $-1 < r < 1$ , then

$$\sum_{k=0}^{\infty} r^k = \frac{1}{1 - r}$$

if the sum is from 1 then

$$\sum_{k=1}^n r^k = \frac{r(1 - r^{n+1})}{1 - r}$$

if  $-1 < r < 1$ , then

$$\sum_{k=1}^{\infty} r^k = \frac{r}{1 - r}$$

$$\Gamma(x) = \int_0^{\infty} u^{x-1} e^{-u} du$$

$$\Gamma(n) = (n - 1)!$$

$$\frac{d}{dx} \ln(x) = \frac{1}{x}$$

$$\int \ln(y) dy = -y + y \ln(y)$$

$$\int \frac{1}{y} dy = \ln(y)$$

$$\binom{n}{n_1 \ n_2 \ n_3} = \frac{n!}{n_1! \ n_2! \ n_3!}$$

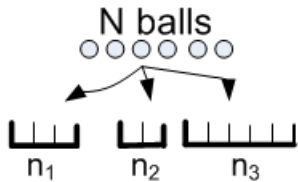
If given joint density  $f_{XY}(x, y)$  and asked to find conditional  $P(X|Y) = \frac{f_{XY}(x, y)}{f_Y(y)}$  so need to find marginals. Marginal is found from  $f_Y(y) = \int_x f_{XY}(x, y) dx$ , and  $f_X(x) = \int_y f_{XY}(x, y) dy$

To convert from  $x, y$  to polar, example: given  $f(x, y) = c\sqrt{1 - (x^2 + y^2)}$  find  $c$ , where  $x^2 + y^2 \leq 1$ , then write

$$c \int_{\theta=-\pi}^{\theta=\pi} \int_{r=0}^{r=1} \sqrt{1 - r^2} r dr d\theta$$

Use identity above.

law of total probability: if we know  $Y|X$  and  $X$  and want to know distribution of  $Y$ , then  $f(Y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) dx$



The Number of WAYS to distribute  $n$  balls into 3 buckets of different sizes is given by the multinomial  $\binom{n}{n_1 \ n_2 \ n_3} = \frac{n!}{n_1! \ n_2! \ n_3!}$

$$r = \int (x + a y)^n dy;$$

$$\text{Out[29]} = \frac{(x + a y)^{1+n}}{a (1 + n)}$$

$$\text{In[2]} := \int r \sqrt{1 - r^2} dr$$

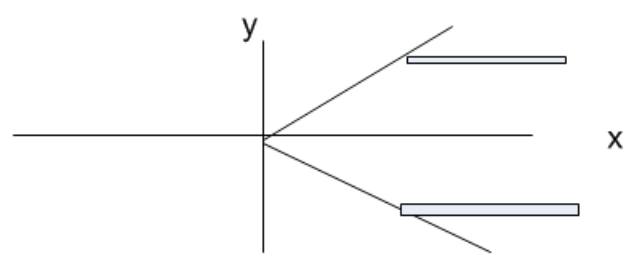
$$\text{Out[2]} = -\frac{1}{3} (1 - r^2)^{3/2}$$

$$\text{In[6]} := \int \sqrt{a^2 - x^2} dx$$

$$\text{Out[6]} = \frac{1}{2} \left( x \sqrt{a^2 - x^2} + a^2 \text{ArcTan} \left[ \frac{x}{\sqrt{a^2 - x^2}} \right] \right)$$

$$\text{In[5]} := \int \sqrt{a^2 + x^2} dx$$

$$\text{Out[5]} = \frac{1}{2} x \sqrt{a^2 + x^2} + \frac{1}{2} a^2 \text{Log} \left[ x + \sqrt{a^2 + x^2} \right]$$



If we need to find marginal density  $f_Y$ , then do it in 2 parts. For  $y > 0, f_Y(y) = \int_{x=-y}^{x=\infty} f(x, y) dx$  and For  $y < 0, f_Y(y) = \int_{x=-y}^{x=\infty} f(x, y) dx$  and leave it at that. do not Add them.

$$Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \rightarrow N(0, 1)$$

$$T = \frac{\bar{X}_n - \mu}{S_n/\sqrt{n}} \rightarrow T(n)$$

where  $S_n$  is *std* of the sample.

Note  $\text{Var}(\text{sample})$  has chi square (n) distribution.

CI for T:

$$\Pr\left(-A < \frac{\bar{X}_n - \mu}{S_n/\sqrt{n}} < A\right) = 1 - \alpha$$
$$\Pr\left(\bar{X}_n - A \frac{S_n}{\sqrt{n}} < \mu < \bar{X}_n + A \frac{S_n}{\sqrt{n}}\right) = 1 - \alpha$$